

## Random Graphs : Markov's Inequality

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## 11.1 Introduction

In this lecture, we will consider the Erdos-Renyi model of random graphs. Our motivation is *not* to present them as a model of graphs that occur in real life—as it is rare to find graphs that behave like Erdos-Renyi graphs. Rather, we present them because they have many counter-intuitive properties, and they provide the most important family of counter-examples to natural conjectures about graphs. Also, their study provides great vehicle for teaching probabilistic analysis.

In this lecture, we will encounter the following quantities associated with graphs.

- Girth. We write  $g(G)$  to denote the girth of a graph  $G$ . It is the length of the *shortest* cycle in  $G$ .
- Clique number. We write  $\omega(G)$  to denote the number of vertices in the largest clique in  $G$ . That is, the largest  $k$  for which there exists a set  $S \subseteq V$  for which all edges between pairs of vertices in  $S$  are in  $G$ .
- Independence number. Written  $\alpha(G)$ , the independence number is the size of the largest set of vertices in  $G$  that has no edges. It is the clique number of the complement graph of  $G$  (the graph that has edges exactly where  $G$  does not).
- Chromatic number, written  $\chi(G)$ . A graph is said to be  $k$ -colorable if there is a mapping  $f : V \rightarrow \{1, \dots, k\}$  so that for every edge  $(u, v)$ ,  $f(u) \neq f(v)$ . The chromatic number of  $G$  is the least  $k$  for which  $G$  is  $k$ -colorable. For example, a bipartite graph is 2-colorable.

Intuitively, one might think that a graph of large girth can be colored with few colors. At the end of lecture, we will see a result of Erdos which tells us this is not true. We will construct the counter-example by the probabilistic method. That is, we will describe a randomized process for constructing a graph, and prove it has the desired properties with non-zero probability. This implies that a graph with the desired properties exists.

## 11.2 Erdos-Renyi Model

The Erdos-Renyi model is specified by two parameters: the number of vertices in the graph  $n$ , and the probability of an edge  $p$ . Given  $n$  and  $p$ , we choose a graph on  $n$  vertices by including an

edge between each pair of vertices with probability  $p$ , independently for each pair. Think of this as flipping a coin for each possible edge. I will write  $\mathcal{G}(n, p)$  to denote this distribution, and

$$G \leftarrow \mathcal{G}(n, p)$$

to indicate that  $G$  is a random graph chosen from this distribution.

### 11.3 Markov's Inequality and Expectation

In this lecture, we will focus on using expectations of random variables. Recall that if a variable  $X$  has the distribution

$$\Pr[X = x_i] = p_i,$$

then

$$\mathbf{E}[X] = \sum_i x_i p_i.$$

The most important property of expectation is that the expectation of the sum of two variables is always the sum of their expectations:

$$\mathbf{E}[X_1 + X_2] = \mathbf{E}[X_1] + \mathbf{E}[X_2].$$

Note that this assertion requires no assumptions! In particular,  $X_1$  and  $X_2$  *do not* need to be independent. This is what makes it so powerful.

If  $X$  is a random variable that can never be negative, then Markov's inequality tells us that for all  $k$

$$\Pr[X \geq k] \leq \mathbf{E}[X] / k.$$

To see why this should be true, note that if the probability that  $X$  is greater than  $k$  is  $p$ , then the expected value of  $X$  would have to be at least  $pk$ .

We will mainly use the following corollary of Markov's inequality:

$$\Pr[X \geq 1] \leq \mathbf{E}[X].$$

### 11.4 Other Facts from Probability

We will also use the *union bound*, which says that for events  $A$  and  $B$ ,

$$\Pr[A \text{ or } B] \leq \Pr[A] + \Pr[B].$$

For  $A$  and  $B$  independent, recall that

$$\Pr[A \text{ and } B] = \Pr[A] \Pr[B].$$

## 11.5 Clique Number

We will now show that for  $p = 1/2$ , the clique number of  $G$  is at most  $(2 + \epsilon) \log_2 n + 1$  with high probability, for all  $\epsilon > 0$ . To do this, fix some  $\epsilon > 0$ , fix  $k = \lceil (2 + \epsilon) \log_2 n + 1 \rceil$ , and let  $S_1, \dots, S_z$  be the subsets of vertices of size  $k$ . So,

$$z = \binom{n}{k}.$$

Let  $X_i$  be a random variable that is 1 if  $S_i$  is a clique in  $G$ . Let

$$X = \sum_i X_i.$$

If  $X < 1$ , then the largest clique in  $G$  has size less than  $k$ . To show that this is probably the case, we will prove that  $\mathbf{E}[X]$  is very small. To do this, we will prove that  $\mathbf{E}[X_i]$  is small for each  $i$ . As  $X_i$  can only take the values 0 and 1,

$$\mathbf{E}[X_i] = \Pr[X_i = 1].$$

We have  $X_i = 1$  only if  $S_i$  is a clique, which happens exactly when all of the  $\binom{k}{2}$  edges between vertices in  $S_i$  appear in the graph. This happens with probability

$$(1/2)^{\binom{k}{2}} = \left( (1/2)^{(k-1)/2} \right)^k = \left( (1/2)^{(2+\epsilon) \log_2 n / 2} \right)^k = \left( (1/2)^{(1+\epsilon/2) \log_2 n} \right)^k = \left( n^{-(1+\epsilon/2)} \right)^k.$$

So,

$$\mathbf{E}[X] = \sum_i \mathbf{E}[X_i] = \binom{n}{k} \left( n^{-(1+\epsilon/2)} \right)^k \leq n^k \left( n^{-(1+\epsilon/2)} \right)^k = \left( n^{-(\epsilon/2)} \right)^k = n^{-\epsilon k / 2} \rightarrow 0.$$

as  $n$  goes to infinity. So, in summary

$$\Pr_{G \leftarrow \mathcal{G}(n, 1/2)} [\omega(G) \geq (2 + \epsilon) \log_2 n + 1] \leq n^{-\epsilon k / 2} \rightarrow 0.$$

We could of course carry this argument out for general  $p$ . This would give

$$\Pr_{G \leftarrow \mathcal{G}(n, p)} [\omega(G) \geq k] \leq \left( np^{(k-1)/2} \right)^k.$$

## 11.6 Girth

In  $\mathcal{G}(n, p)$ , the expected number of edges attached to each vertex is  $p(n-1)$ . Next lecture, we will see that it must be close to this for all vertices. For now, let's set  $d = p(n-1)$ , and ask ourselves how large the girth can be of a graph in which every vertex has degree  $d$ . We will then show that this bound is almost achieved.

For a vertex  $v$ , let  $N(v)$  be the set of neighbors of  $v$ , and let  $N^{(k)}(v)$  be the set of vertices that can be reached from  $v$  by a path of length at most  $k$ . We have  $|N(v)| = d + 1$ . If  $g(G) > 4$ , then

$$|N^{(2)}(v)| = d(d-1) + |N(v)| = d^2 + 1.$$

Similarly, if  $g(G) \geq 2k + 1$ , then

$$|N^{(k)}(v)| = d(d-1)^k + |N^{(k-1)}(v)| = d^k + 1.$$

But, there are  $n$  vertices, so

$$n \geq |N^{(k)}(v)| = d^k + 1,$$

which implies

$$k < \log_d n = \frac{\log n}{\log d}.$$

If  $d = n^{1/j}$ , this gives the bound  $g \leq 2j + 1$ . We will now show that this is approximately tight.

Set  $p = n^{(1-\epsilon)/g}/n$ , for any  $\epsilon > 0$ , and choose  $G \leftarrow \mathcal{G}(n, p)$ . We will prove that few of the vertices of  $G$  will be in cycles of length  $g$ . It would be unreasonable to hope that there are no short cycles.

(In fact, the analysis from the previous section tells us that the expected number of triangles is  $\binom{n}{3}p^3 = n^{3(1-\epsilon)/g}/6 > 1$ .)

A  $g$ -cycle is described by a sequence of  $g$  vertices, giving the first vertex in the cycle, the second, and so on. Actually, each  $g$ -cycle has  $2g$  descriptions of this form: there are  $g$  choices for the first vertex, and two directions in which the cycle can be traversed. Either way, we know that there are most

$$n(n-1) \cdots (n-g+1) \leq n^g$$

possible  $g$ -cycles. The probability that any given possible  $g$ -cycle appears in  $G$  is  $p^g$ . So, the expected number of  $g$ -cycles is at most

$$n^g p^g = (np)^g = \left(n^{(1-\epsilon)/g}\right)^g = n^{1-\epsilon}.$$

One can show that the expected number of  $j$  cycles for  $j < g$  is lower. So, the expected number of cycles of length at most  $g$  is at most

$$gn^{1-\epsilon}.$$

By Markov's inequality, this means that the probability that  $G$  has more than  $2gn^{1-\epsilon}$  cycles of length at most  $g$  is at most  $1/2$ , and that the probability that  $G$  has more than  $n/2g$  cycles of length up to  $g$  is at most

$$gn^{1-\epsilon}/(n/2g) = 2g^2/n^\epsilon.$$

So, we may conclude that the probability that more than  $n/2$  of the vertices are involved in cycles of length up to  $g$  is at most

$$2g^2/n^\epsilon.$$

Note that all of these bounds go to zero as  $n$  grows and  $g$  stays fixed.

Consider removing all the vertices in  $G$  that are involved in cycles of length up to  $g$ . With probability  $1 - 2g^2/n^\epsilon$ , at least  $n/2$  vertices remain, and the remaining graph has girth at least  $g$ . You might be wondering how many edges are left in the graph. We will later learn techniques that show that with high probability at least  $1/4$  of the edges remain.

## 11.7 High Girth and Chromatic Number

**Theorem 11.7.1 (Erdos).** *For every  $g$  and  $x$ , there exists a graph  $G$  of girth at least  $g$  and chromatic number at least  $x$ .*

*Proof.* As in the previous section, set  $p = n^{(1-\epsilon)/g}/n$ , and choose  $G$  from  $\mathcal{G}(n, p)$ . Use  $\epsilon = 1/2$ , so  $p = n^{1/2g}/n$ .

Then, remove all vertices from  $G$  in cycles of length up to  $g$ , and call the resulting graph  $G'$ . We will show that  $G'$  probably has high chromatic number.

We would like to say something like “ $\chi(G') \geq \chi(G)$ ”, but I see no reason it should be true. Instead, we use the inequalities

$$\alpha(G) \geq \frac{n}{\chi(G)}, \quad (11.1)$$

and

$$\alpha(G') \leq \alpha(G). \quad (11.2)$$

The first follows from the fact that each color class in a coloring is an independent set, and the largest must have size at least  $n/\chi(G)$ . The second is because every independent set in  $G'$  is also an independent set in  $G$ . If we let  $n'$  be the number of vertices in  $G'$ , we may combine these inequalities to find

$$\chi(G') \geq \frac{n'}{\alpha(G')} \geq \frac{n'}{\alpha(G)}.$$

Let's see what we can say about  $\alpha(G)$ . As  $\alpha(G)$  is the clique number of the complement graph of  $G$ , we may apply the results from the first section to count the expected number of independent set in  $G$  of size  $a$ . Let  $X$  be the number of independent set in  $G$  of size  $a$ . We get

$$\mathbf{E}[X] \leq \left( n(1-p)^{(a-1)/2} \right)^a.$$

To estimate this, we will use one of the most important inequalities in probability:

$$1 - p \leq e^{-p}.$$

I suggest you memorize it.

We then compute

$$\begin{aligned} \mathbf{E}[X] &\leq \left( n(1-p)^{(a-1)/2} \right)^a \\ &\leq \left( n(e^{-p})^{(a-1)/2} \right)^a \\ &= \left( n(e^{-p(a-1)/2}) \right)^a. \end{aligned}$$

If we set

$$a = \frac{4n \ln n}{n^{1/2g}} + 1,$$

then we get

$$\left(n(e^{-p(a-1)/2})\right)^a = \left(n(e^{-2\ln n})\right)^a = (nn^{-2})^a = n^{-a}.$$

So, the probability that  $\alpha(G)$  exceeds  $a$  is at most  $n^{-a}$ . Now, if  $\alpha(G) \leq a$  and  $n' \geq n/2$ , then

$$\chi(G') \geq \frac{n'}{\alpha(G)} \geq \frac{n/2}{\frac{4n \ln n}{n^{1/2g}} + 1} = \frac{n^{1/2g}}{8 \ln n + n^{1/2g-1}} \geq \frac{n^{1/2g}}{9 \ln n}.$$

If we fix  $g$  and let  $n$  grow, then this quantity grows as well, and so eventually becomes bigger than  $x$ . So, to establish the existence of the desired graph  $G'$ , we just need to show that with some reasonable probability,  $n' \geq n/2$  and  $\alpha(G) \leq a$ . To do this, we examine the probability of failing. We have

$$\Pr [n' \leq n/2 \text{ or } \alpha(G) > a] \leq \Pr [n' \leq n/2] + \Pr [\alpha(G) > a] \leq 2g^2/n^\epsilon + n^{-a} \rightarrow 0,$$

as  $n$  grows. So, the probability of  $G'$  having the desired properties tends to 1, and so the desired graph exists.

In fact, we only needed to show that the probability of  $G'$  having the desired properties is greater than 0 □